

**INVESTIGATIONS ON THE SATURATION OF C.D.F. OF  
THE MODIFIED "ODD GENERALIZED EXPONENTIAL  
POWER FUNCTION DISTRIBUTION". SOME APPLICATIONS**

TODORKA TERZIEVA<sup>1</sup>, ANTON ILIEV<sup>2</sup>, ASEN RAHNEV<sup>3</sup>,  
AND NIKOLAY KYURKCHIEV<sup>4</sup>

<sup>1,2,3,4</sup>Faculty of Mathematics and Informatics  
University of Plovdiv Paisii Hilendarski  
24, Tzar Asen Str., 4000 Plovdiv, BULGARIA

**ABSTRACT:** In [1] the authors proposed a new family of continuous distributions called the odd generalized exponential family.

Their reviews have brought to life a number of modifications that have appeared in the literature over the last few years, with applications to approximate specific "cumulative data" from different branches of science.

In this article we will look at the one suggested in [2].

The authors' assertion that probability distribution produces very good results encouraged us to conduct further studies on "saturation" in Hausdorff sense of the corresponding commutative function to the horizontal asymptote hoping to partially contribute to uncovering some of the "intrinsic properties" of this apparently good model.

We also analyze some experimental data: the cumulative number of Welchia attackers; "growth of the cumulative number of TREZ publications"; data set "US mobile"; Software Failure Data – Release #1; Defects (AT & T).

Numerical examples, illustrating our results are presented using programming environment *CAS Mathematica*.

**AMS Subject Classification:** 41A46

**Key Words:** c.d.f. of the ODD–Generalized–Exponential–Power–Function–Distribution (cdfOGEPPF), Heaviside step–function  $h_{t_0}(t)$ , Hausdorff distance

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## 1. INTRODUCTION AND PRELIMINARIES

In this note we study the Hausdorff approximation of the Heaviside function  $h_{t_0}(t)$  by the cdf of the ODD–Generalized–Exponential–Power–Function–Distribution (cdfOGEFP), suggested in [2].

The model have been tested with real-world data.

**Definition 1.** The cdf of the ODD–Generalized–Exponential–Power–Function–Distribution (cdfOGEFP) is defined by [2]:

$$M(t) = \left(1 - e^{-\lambda\left(\frac{t^\alpha}{\beta^\alpha - t^\alpha}\right)}\right)^\theta \quad (1)$$

where  $0 \leq t < \beta$ ,  $\theta > 0$ ,  $\alpha > 0$ .

**Definition 2.** The shifted Heaviside step function is defined by

$$h_{t_0}(t) = \begin{cases} 0, & \text{if } t < t_0, \\ [0, 1], & \text{if } t = t_0, \\ 1, & \text{if } t > t_0 \end{cases} \quad (2)$$

**Definition 3.** [3] The Hausdorff distance (the H–distance)  $\rho(f, g)$  between two interval functions  $f, g$  on  $\Omega \subseteq \mathbb{R}$ , is the distance between their completed graphs  $F(f)$  and  $F(g)$  considered as closed subsets of  $\Omega \times \mathbb{R}$ . More precisely,

$$\rho(f, g) = \max\left\{ \sup_{A \in F(f)} \inf_{B \in F(g)} \|A - B\|, \sup_{B \in F(g)} \inf_{A \in F(f)} \|A - B\| \right\},$$

wherein  $\|\cdot\|$  is any norm in  $\mathbb{R}^2$ , e. g. the maximum norm  $\|(t, x)\| = \max\{|t|, |x|\}$ ; hence the distance between the points  $A = (t_A, x_A)$ ,  $B = (t_B, x_B)$  in  $\mathbb{R}^2$  is  $\|A - B\| = \max(|t_A - t_B|, |x_A - x_B|)$ .

## 2. MAIN RESULTS

### 2.1. A NOTE ON THE NEW (CDFOGEFP)

The investigation of the characteristic "supersaturation" of the cdf (1) to the horizontal asymptote is important.

The quantile function is defined by:

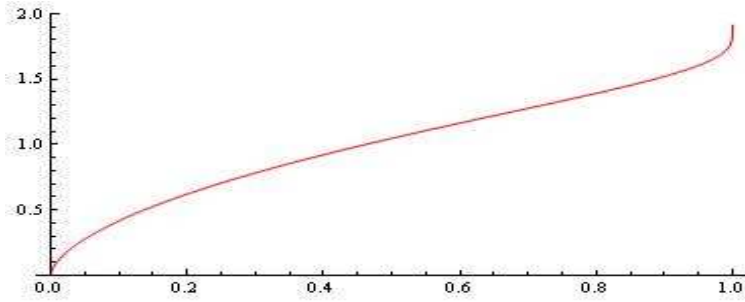


Figure 1: The function  $Q(u)$  for  $\alpha = 1.9, \beta = 2, \lambda = 1.5$  and  $\theta = 0.9$ .

$$Q(u) = \beta \left( -\frac{\ln(1 - u^{\frac{1}{\theta}})}{\lambda - \ln(1 - u^{\frac{1}{\theta}})} \right)^{\frac{1}{\alpha}} \tag{3}$$

(see, Fig. 1 for  $\alpha = 1.9, \beta = 2, \lambda = 1.5$  and  $\theta = 0.9$ ).

The median is obtained by substituting  $u = 0.5$  in (3).

Let  $t_0$  is the value for which  $M(t_0) = \frac{1}{2}$ , i.e.  $t_0 = u(0.5)$  .

The one-sided Hausdorff distance  $d$  between the function  $h_{t_0}(t)$  and the cdf (1) satisfies the relation

$$M(t_0 + d) = 1 - d. \tag{4}$$

For given  $\beta, \theta, \alpha, \lambda$ , the nonlinear equation  $M(t_0 + d) - 1 + d = 0$  has unique positive root  $-d$ .

The cdf (1) for  $\lambda = 1.5, \beta = 2, \theta = 0.9, \alpha = 1.9$  and  $t_0 = 1.04818$  is visualized on Fig. 2.

From the nonlinear equation (4) we have:  $d = 0.266933$ .

The cdf (1) for  $\lambda = 25.6, \beta = 1.8, \theta = 0.999, \alpha = 1.4$  and  $t_0 = 0.133999$  is visualized on Fig. 3.

From the nonlinear equation (4) we have:  $d = 0.139343$ .

From the above examples, it can be seen that the "supersaturation" by the (cdf)  $M(t)$  is faster.

Obviously, this "advantage" can actually be used to approximate some specific "cumulative data" from different branches of science.

In the next Section, we will support what is said by analyzing real datasets.

## 2.2. APPLICATIONS

Welchia worm ransomware has a long growing phase in contrast to many other threats.

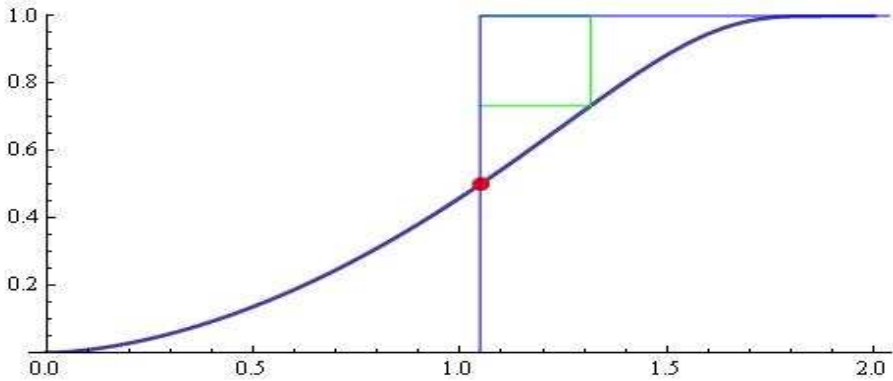


Figure 2: The cdf (1) for  $\lambda = 1.5$ ,  $\beta = 2$ ,  $\theta = 0.9$ ,  $\alpha = 1.9$  and  $t_0 = 1.04818$ ; H-distance  $d = 0.266933$ .

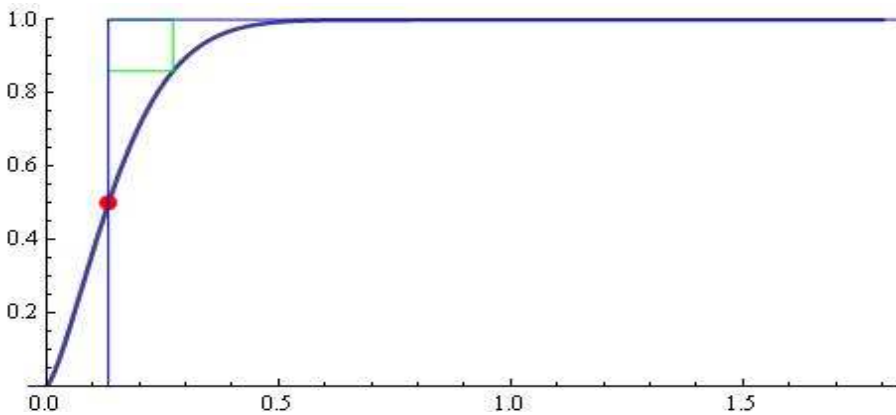


Figure 3: The cdf (1) for  $\lambda = 25.6$ ,  $\beta = 1.8$ ,  $\theta = 0.999$ ,  $\alpha = 1.4$  and  $t_0 = 0.133999$ ; H-distance  $d = 0.139343$ .

Welchia worm uses a vulnerability in the Microsoft remote procedure call service. Welchia firstly checks for Blaster worm and if it exists continues with Blaster deletion as well as takes care for computer to be immunised for Blaster worm.

**Example 1.** Analysis of Welchia worm infection behavior

For epidemic as Welchia worm it is appropriately to use a model

$$M^*(t) = \omega \left( 1 - e^{-\lambda \left( \frac{t^\alpha}{\beta^\alpha - t^\alpha} \right)} \right)^\theta \quad (5)$$

for approximating data from the statistics collected on an individual Welchia [4] non-epidemic administered by Frederic Perriot between August 24th, 2003 and February 24th,

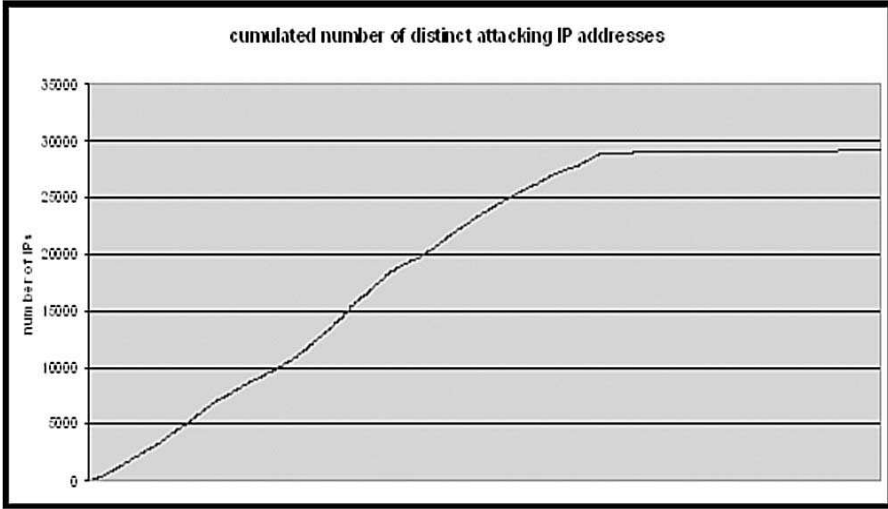


Figure 4: The cumulative number of Welchia attackers [4].

2004, shown in Fig. 4.

We will explore this example by photographing the data from Fig. 4.

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data_Welchia :=
    {{1.1, 1000}, {2, 2333}, {3, 3500}, {4, 5000}, {5, 6833}, {6, 8000},
    {7, 9333}, {8, 10500}, {9, 12000}, {10, 14000}, {11, 16333},
    {12, 18167}, {13, 19667}, {14, 21000}, {15, 22667}, {16, 23667},
    {17, 25000}, {18, 26333}, {19, 27500}, {20, 28333}, {21, 29333},
    {22, 29500}, {23, 29500}, {24, 29500}, {25, 29500}, {26, 29500},
    {27, 29500}, {28, 29500}, {29, 29500}, {30, 29500}, {31, 29667},
    {32, 29667}}
    
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The fitted model is given by

$$\lambda = 2.51375; \beta = 33; \theta = 0.999; \alpha = 1.31223; \omega = 30000.$$

We receive an impressive result when approximating these data, see Fig. 5.

Comparison between  $M^*(t)$  and the model  $M_1^*(t)$  [5] with  $\lambda = -0.1; \beta = 0.0385232; \theta = 0.602142; \omega = 30000$  (see, Fig. 6):

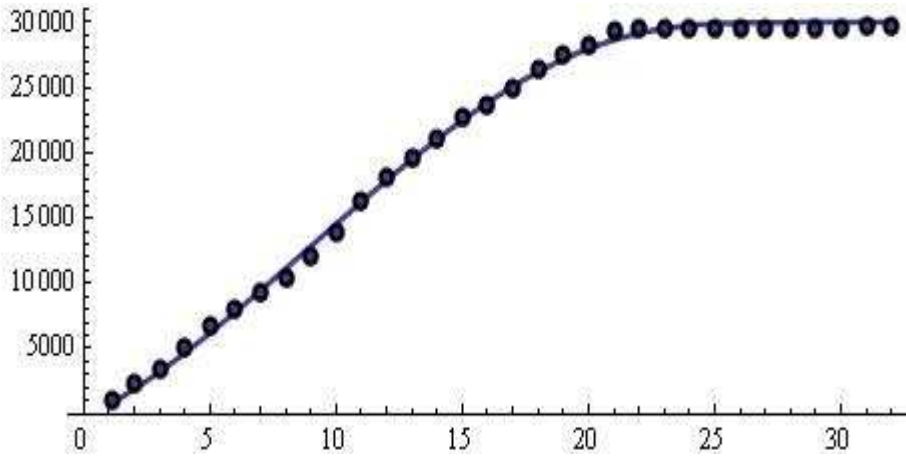


Figure 5: The fitted model  $M^*(t)$ .

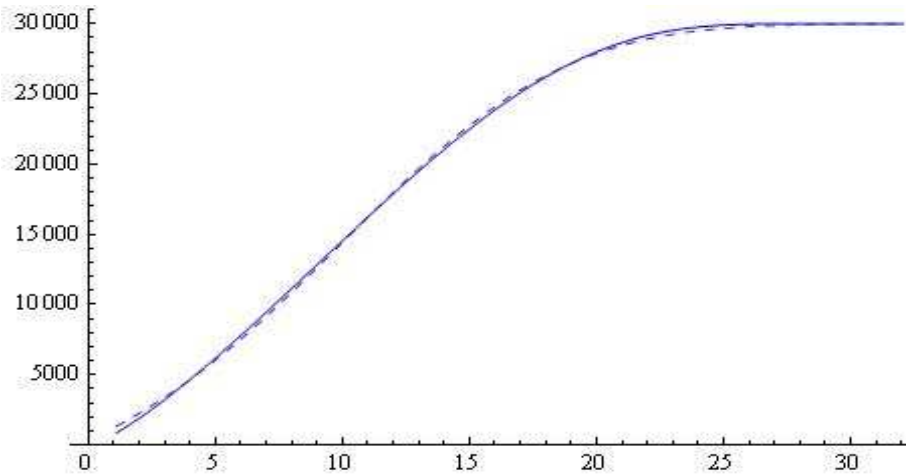
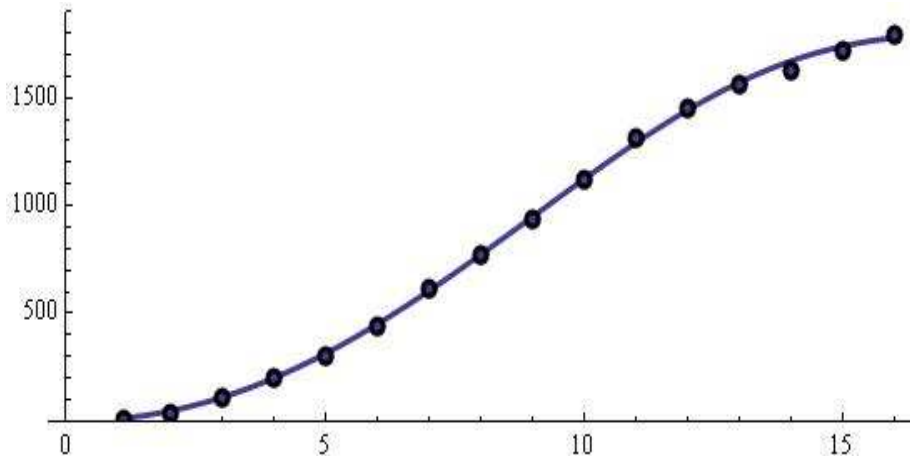


Figure 6: The fitted models  $M^*(t)$ —blue and  $M_1^*(t)$ —dashed.

$$M_1^*(t) = \omega \left( 1 - \left( 1 + \frac{e^{\theta((1-\lambda t^\theta)^{-\frac{1}{\lambda}} - 1)} - 1}{e^\theta - 1} \right)^{-\beta} \right) \tag{6}$$

**Example 2.** Analysis of data "growth of the cumulative number of TREZ publications" [6], [7]

Figure 7: The fitted model  $M^*(t)$ .

*data\_Journal*

:= {{1.1, 5}, {2, 37}, {3, 107}, {4, 201}, {5, 298}, {6, 439},  
 {7, 617}, {8, 773}, {9, 936}, {10, 1121}, {11, 1316},  
 {12, 1451}, {13, 1563}, {14, 1629}, {15, 1722}, {16, 1788}};

After that using the model  $M^*(t) = \omega M(t)$  for  $\lambda = 4.80604$ ,  $\beta = 22.1$ ,  $\theta = 0.9$ ,  $\alpha = 2.33693$  and  $\omega = 1800$  we obtain the fitted model (see, Fig. 7).

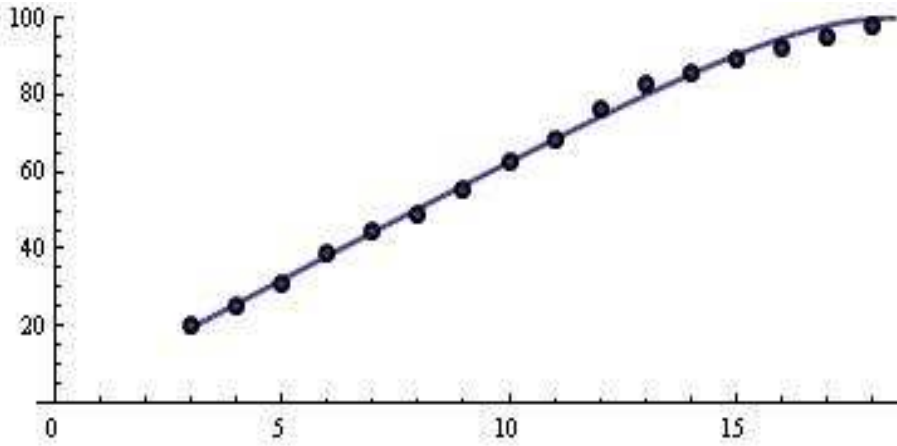
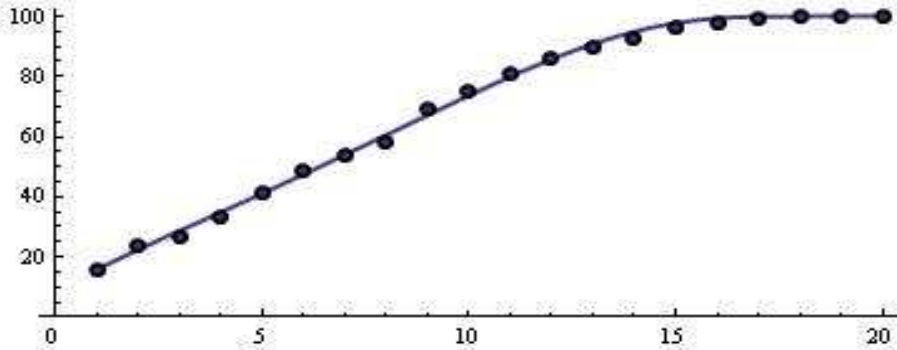
**Example 3.** Analysis of data set "US mobile", [7]

*dataMobile*

:= {{3, 20.29}, {4, 25.08}, {5, 30.81}, {6, 38.75}, {7, 45}, {8, 49.16},  
 {9, 55.15}, {10, 62.852}, {11, 68.63}, {12, 76.64}, {13, 82.47},  
 {14, 85.68}, {15, 89.14}, {16, 91.864}, {17, 95.28}, {18, 98.17}};

After that using the model  $M^*(t) = \omega M(t)$  for  $\lambda = 3.10703$ ,  $\beta = 20.1$ ,  $\theta = 0.183849$ ,  $\alpha = 5.26477$  and  $\omega = 100$  we obtain the fitted model (see, Fig. 8).

**Example 4.** Software Failure Data – Release #1. These data derive from one major release of software products at Tandem Computers [8], [9]

Figure 8: The fitted model  $M^*(t)$ .Figure 9: The fitted model  $M^*(t)$ .
$$\begin{aligned} \text{data\_PhamNordmanZhang} := & \{\{1, 16\}, \{2, 24\}, \{3, 27\}, \{4, 33\}, \\ & \{5, 41\}, \{6, 49\}, \{7, 54\}, \{8, 58\}, \{9, 69\}, \{10, 75\}, \{11, 81\}, \{12, 86\}, \\ & \{13, 90\}, \{14, 93\}, \{15, 96\}, \{16, 98\}, \{17, 99\}, \{18, 100\}, \{19, 100\}, \\ & \{20, 100\}\}; \end{aligned}$$

For this data the fitted model for estimated parameters:  $\omega = 100$ ;  $\alpha = 0.0265009$ ;  $\beta = 20.1$ ;  $\lambda = 0.0336371$ ;  $\theta = 1.69507$  is plotted on Fig. 9.



### 3. CONCLUDING REMARKS

Finally, we note that the studied model produces extremely good results, generally when approximating specific "cumulative data" from Computer Viruses Propagation, Debugging and Test theory.

For other approximation and modelling results, see [10]–[23].

We hope that the results will be useful for specialists in this scientific area.

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